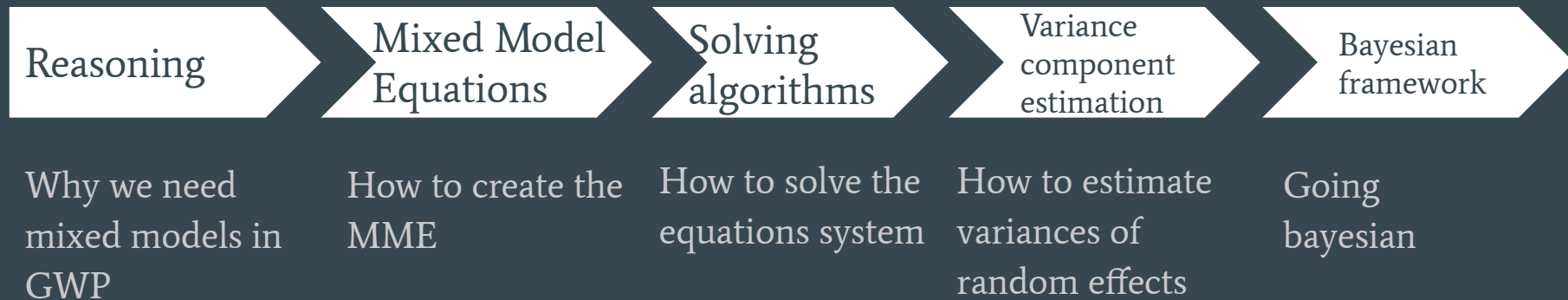


# Linear Mixed Models

...

Evangelina López de Maturana & Oscar González-Recio

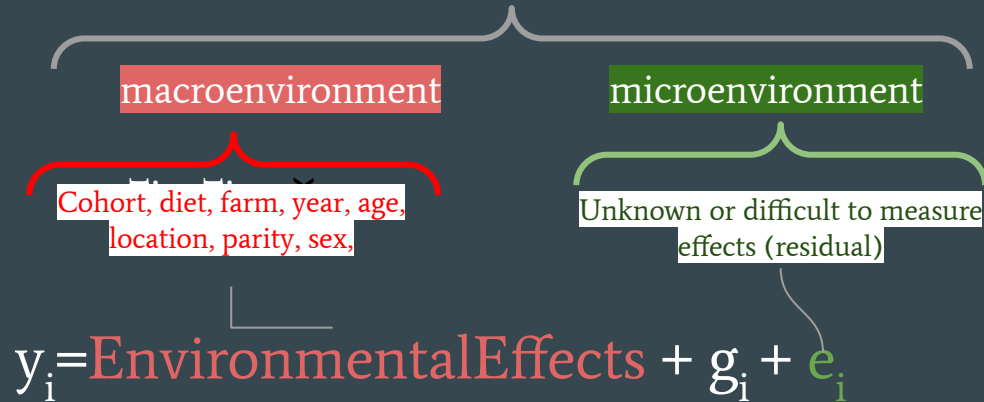
# What you need to know



# Mathematical representation of biological processes

$$P=G+E$$

$$y_i = g_i + E_i$$



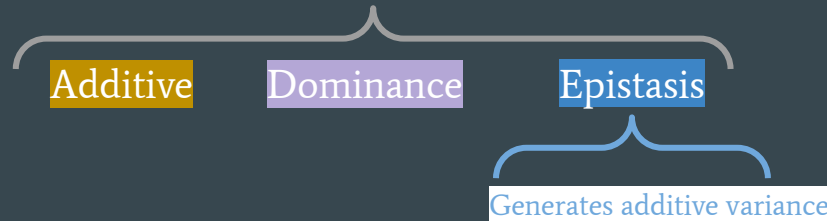
$$y_i = X_i b_i + Z_i g_i + e_i$$



# Mathematical representation of biological processes

$$P=G+E$$

$$y_i = g_i + E_i$$



Hill et al (2008) <https://doi.org/10.1371/journal.pgen.1000008>



$$y_i = X_i b_i + Z_{ui} u_i + Z_{di} d_i + Z_{ei} u_i \# u_i + Z_{ei} u_i \# d_i + \dots + e_i$$

$$y_i = X_i b_i + Z_{ui} u_i + e'_i$$

# Mathematical representation of biological processes

$$P=G+E$$

$$y_i = g_i + E_i$$

Additive

Dominance

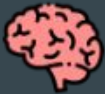
Epistasis

Generates additive variance

Hill et al (2008) <https://doi.org/10.1371/journal.pgen.1000008>

$$y_i = X_i b_i + Z_{ui} u_i + Z_{di} d_i + Z_{ei} u_i \# u_i + Z_{ei} u_i \# d_i + \dots + e_i$$

$$y_i = X_i b_i + Z_{ui} u_i + e'_i$$



# Linear mixed models

$$y_i = X_i b_i + Z_{ui} u_i + e'_i$$

$$\mathbf{e} \sim N(0, \mathbf{I} \otimes \sigma_e^2)$$

$$\mathbf{u} \sim N(0, \mathbf{A} \otimes \sigma_a^2)$$

“Fixed” terms

“Random” terms (assume some known distribution to the effect)



Genome-wide prediction



# Linear mixed models: Mixed Model Equation (EMM)

$$y_i = X_i b_i + Z_{ui} u_i + e'_i$$



$$\begin{pmatrix} Y_{111} \\ Y_{112} \\ Y_{113} \\ Y_{114} \\ Y_{115} \\ Y_{121} \\ Y_{122} \\ Y_{123} \\ Y_{124} \\ Y_{125} \\ Y_{211} \\ Y_{212} \\ Y_{213} \\ Y_{214} \\ Y_{215} \\ Y_{221} \\ Y_{222} \\ Y_{223} \\ Y_{224} \\ Y_{225} \\ Y_{311} \\ Y_{312} \\ Y_{313} \\ Y_{314} \\ Y_{315} \\ Y_{321} \\ Y_{322} \\ Y_{323} \\ Y_{324} \\ Y_{325} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \end{pmatrix} \begin{pmatrix} \mu \\ \tau_1 \\ \tau_2 \end{pmatrix} + \begin{pmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} h_1 \\ h_2 \\ h_3 \end{pmatrix} + \begin{pmatrix} e_{111} \\ e_{112} \\ e_{113} \\ e_{114} \\ e_{115} \\ e_{121} \\ e_{122} \\ e_{123} \\ e_{124} \\ e_{125} \\ e_{211} \\ e_{212} \\ e_{213} \\ e_{214} \\ e_{215} \\ e_{221} \\ e_{222} \\ e_{223} \\ e_{224} \\ e_{225} \\ e_{311} \\ e_{312} \\ e_{313} \\ e_{314} \\ e_{315} \\ e_{321} \\ e_{322} \\ e_{323} \\ e_{324} \\ e_{325} \end{pmatrix}$$

$Y = X \beta + Z u + e$

Example of  
fixed effects  
(mean, and an  
effect with 2  
levels)

Example of  
random effect  
with 3 levels  
(ignoring the  
covariance  
structure)

# Linear mixed models: Mixed Model Equation (MME)

$$y_i = \mathbf{X}_i \mathbf{b}_i + \mathbf{Z}_{ui} \mathbf{u}_i + \mathbf{e}'_i$$

When  $\mathbf{X}$  and  $\mathbf{Z}$  are too large, inversion of matrices is very computationally demanding, there are linear combinations, and solving the system becomes cumbersome.

Henderson proposed MME

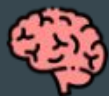
$$\begin{bmatrix} \mathbf{X}'\mathbf{X} & \mathbf{X}'\mathbf{Z} \\ \mathbf{Z}'\mathbf{X} & \mathbf{Z}'\mathbf{Z} + \alpha \mathbf{A}^{-1} \end{bmatrix} \begin{bmatrix} \hat{\boldsymbol{\beta}} \\ \hat{\mathbf{u}} \end{bmatrix} = \begin{bmatrix} \mathbf{X}'\mathbf{Y} \\ \mathbf{Z}'\mathbf{Y} \end{bmatrix}$$



# Linear mixed models: Mixed Model Equation (EMM)

Increased complexity

$$\mathbf{y}_i = \mathbf{X}_i \mathbf{b}_i + \mathbf{U}_i \mathbf{u}_i + \mathbf{W}_i \mathbf{p}_i + \mathbf{Z}_{ui} \mathbf{u}_i + \mathbf{e}'_i$$



$$\begin{pmatrix} \mathbf{X}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{X}'\mathbf{R}^{-1}\mathbf{U} & \mathbf{X}'\mathbf{R}^{-1}\mathbf{W} & \mathbf{X}'\mathbf{R}^{-1}\mathbf{Z} \\ \mathbf{U}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{U}'\mathbf{R}^{-1}\mathbf{U} + \Theta_1 & \mathbf{U}'\mathbf{R}^{-1}\mathbf{W} & \mathbf{U}'\mathbf{R}^{-1}\mathbf{Z} \\ \mathbf{W}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{W}'\mathbf{R}^{-1}\mathbf{U} & \mathbf{W}'\mathbf{R}^{-1}\mathbf{W} + \Theta_2 & \mathbf{W}'\mathbf{R}^{-1}\mathbf{Z} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{U} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{W} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \Theta_S \end{pmatrix}^{-1} \mathbf{X} \begin{pmatrix} \mathbf{X}'\mathbf{R}^{-1}\mathbf{Y} \\ \mathbf{U}'\mathbf{R}^{-1}\mathbf{Y} \\ \mathbf{W}'\mathbf{R}^{-1}\mathbf{Y} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Y} \end{pmatrix}$$



Genome-wide prediction



# Linear mixed models: Mixed Model Equation (EMM)

Marker regression

Decomposing the polygenic effect into the sum of SNP effects

$$y_i = \mathbf{X}_i \mathbf{b}_i + \mathbf{Z}_{ui} \mathbf{u}_i + \mathbf{e}'_i$$

$$\mathbf{y} = \mu \mathbf{1} + \mathbf{X} \boldsymbol{\beta}_f + \textit{snp}_1 \beta_1 + \textit{snp}_2 \beta_2 + \dots + \textit{snp}_p \beta_p + \mathbf{e}$$

$$\mathbf{e} \sim N(0, \sigma_e^2)$$

$$\mathbf{b}_i \sim N(0, \sigma_i^2)$$

$$\sigma_i^2 \sim \chi_{(v,S)}^{-2}$$



Genome-wide prediction



# Variance component estimation

## Methods

Estimate residual, genetic, permanent, marker, ... variances

- ANOVA
- Maximum Likelihood (ML)
- Restricted Maximum Likelihood (REML)
- Minimum Norm Quadratic Unbiased Estimation (MINQUE I, II, III)
- Minimum Variance Quadratic Unbiased Estimation (MIVQUE)



### Journal of the American Statistical Association

Publication details, including instructions for authors and subscription information:

<http://www.tandfonline.com/loi/uasa20>

### A Comparison of Variance Component Estimates for Arbitrary Underlying Distributions

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Published online: 12 Mar 2012.

# Solve MME

## Algorithms

- Gauss-seidel
- Choleski decomposition
- Preconditioned conjugate gradients (PCG)
- Gauss-seidel with residual updates

J. Dairy Sci. 91:360–366  
doi:10.3168/jds.2007-0403

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### ***Technical Note:*** Computing Strategies in Genome-Wide Selection

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# Solve MME

## Algorithms

2,000 records  
11,000 SNPs

**Table 1.** Computing times for different methods of solving the mixed-model equations in a case of genome-wide genetic evaluation<sup>1</sup>

Item	Cholesky decomposition <sup>2</sup>	LHS-GS	Matrix-free GS	Matrix-free GSRU	Matrix-free PCG
Convergence = $10^{-10}$					
Time to set up	17 min	17 min	16 s	16 s	20 s
Time for solving	119 min	71 min	97 h	46 s	10 s
Number of iterations	1	164	161	164	20
Convergence = $10^{-14}$					
Time to set up	17 min	17 min	<sup>3</sup>	16 s	20 s
Time for solving	119 min	170 min	<sup>3</sup>	74 s	12 s
Number of iterations	1	272	<sup>3</sup>	272	23

<sup>1</sup>LHS = left-hand side; GS = Gauss-Seidel; GSRU = Gauss-Seidel with residual update; PCG = preconditioned conjugated gradients.

<sup>2</sup>This is an exact method and the convergence measure is meaningless.

<sup>3</sup>Not tried.

# Solve MME

## Gauss Seidel with Residual Update

$y$  corrected for all effects except the  $j^{\text{th}}$  effect, is equal to the current vector of residuals plus the current estimate of the  $j^{\text{th}}$  effect. Then, we can compute this  $j^{\text{th}}$  effect.

$$y - \mathbf{X}_{1:j-1} \hat{\mathbf{a}}_{1:j-1}^{l+1} - \mathbf{X}_{j+1:n} \hat{\mathbf{a}}_{j+1:n}^l = \mathbf{e}^{l+1,j} + \mathbf{x}_j \hat{a}_j^l.$$

Then

$$\hat{a}_j^{l+1} = \frac{\mathbf{x}_j' \mathbf{e}^{l+1,j} + \mathbf{x}_j' \mathbf{x}_j \hat{a}_j^l}{\mathbf{x}_j' \mathbf{x}_j + \lambda}.$$

We update the residuals with the new  $j^{\text{th}}$  effect solution.

$$\mathbf{e}^{l+1,j+1} = \mathbf{e}^{l+1,j} - \mathbf{x}_j (\hat{a}_j^{l+1} - \hat{a}_j^l).$$

Solution from previous iteration

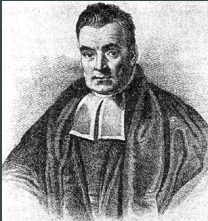
Updated solution in the new iteration

$\mathbf{x}_j' \mathbf{x}_j$  are constant, so they can be precomputed, updating vector products and residuals at each iteration, speeding up the algorithm

# Bayesian framework

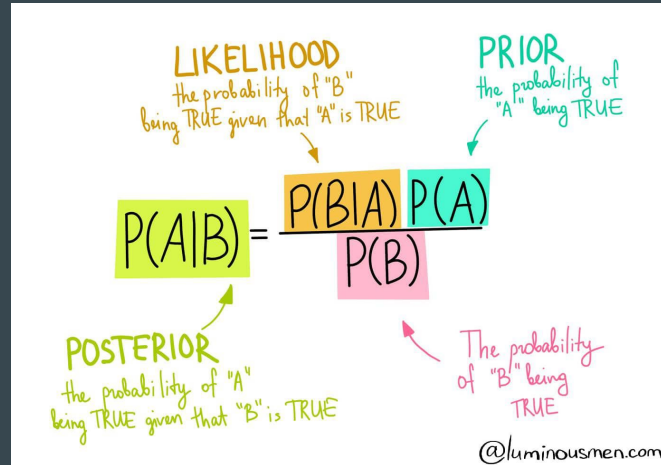
Assume a model for the data

Bayes theorem



$$\mathbf{y} = \mu \mathbf{1} + \mathbf{X}\boldsymbol{\beta}_f + \text{sn}p_1\boldsymbol{\beta}_1 + \text{sn}p_2\boldsymbol{\beta}_2 + \dots + \text{sn}p_p\boldsymbol{\beta}_p + \mathbf{e}$$

$$p(\mathbf{y}|\boldsymbol{\theta}) = N(\mathbf{X}\boldsymbol{\beta} + \dots + \mathbf{Z}\mathbf{u}, \sigma_e^2)$$

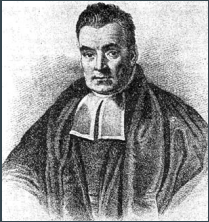


# Bayesian framework

Assume a model for the data

$$p(\mathbf{y}|\boldsymbol{\theta}) = N(\mathbf{X}\boldsymbol{\beta} + \cdots + \mathbf{Z}\mathbf{u}, \sigma_e^2)$$

Bayes theorem



$$p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta}) = p(\boldsymbol{\theta}|\mathbf{y})p(\mathbf{y})$$

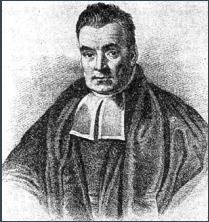


# Bayesian framework

Assume a model for the data

$$p(\mathbf{y}|\boldsymbol{\theta}) = N(\mathbf{X}\boldsymbol{\beta} + \dots + \mathbf{Z}\mathbf{u}, \sigma_e^2)$$

Bayes theorem



$$p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta}) = p(\boldsymbol{\theta}|\mathbf{y})p(\mathbf{y})$$

...proportionality

$$p(\boldsymbol{\theta}|\mathbf{y}) = \frac{p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{p(\mathbf{y})} \propto p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta})$$

$\mathbf{y}$ =data

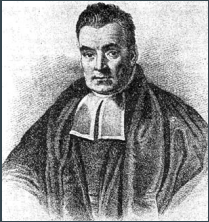
$\boldsymbol{\theta}$ = unknown parameters, coefficients, variances, ...

# Bayesian framework

Assume a model for the data

$$p(\mathbf{y}|\boldsymbol{\theta}) = N(\mathbf{X}\boldsymbol{\beta} + \dots + \mathbf{Z}\mathbf{u}, \sigma_e^2)$$

Bayes theorem



$$p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta}) = p(\boldsymbol{\theta}|\mathbf{y})p(\mathbf{y})$$

...proportionality

$$p(\boldsymbol{\theta}|\mathbf{y}) = \frac{p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{p(\mathbf{y})} \propto p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta})$$

Choose priors

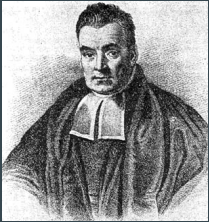
$$p(\boldsymbol{\theta}|\mathbf{y}) \propto p(\mathbf{y}|\boldsymbol{\beta}, \mathbf{u}, \sigma_e^2)p(\boldsymbol{\beta}|\sigma_b^2)p(\mathbf{u}|\sigma_u^2)p(\sigma_b^2)p(\sigma_u^2)p(\sigma_e^2)$$

# Bayesian framework

Assume a model for the data

$$p(\mathbf{y}|\boldsymbol{\theta}) = N(\mathbf{X}\boldsymbol{\beta} + \dots + \mathbf{Z}\mathbf{u}, \sigma_e^2)$$

Bayes theorem



$$p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta}) = p(\boldsymbol{\theta}|\mathbf{y})p(\mathbf{y})$$

...proportionality

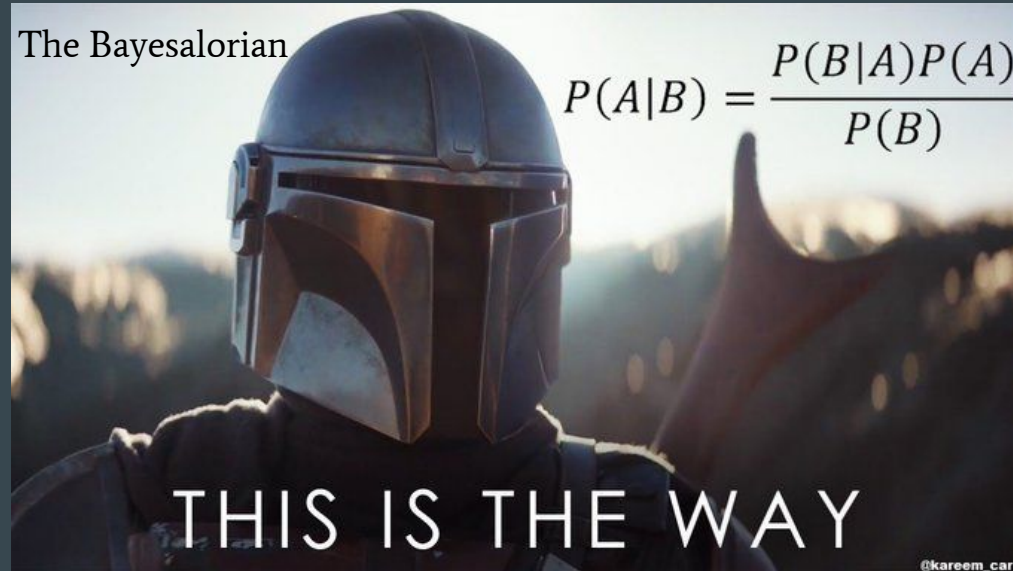
$$p(\boldsymbol{\theta}|\mathbf{y}) = \frac{p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{p(\mathbf{y})} \propto p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta})$$

Choose priors

$$p(\boldsymbol{\theta}|\mathbf{y}) \propto p(\mathbf{y}|\boldsymbol{\beta}, \mathbf{u}, \sigma_e^2)p(\boldsymbol{\beta}|\sigma_b^2)p(\mathbf{u}|\sigma_u^2)p(\sigma_b^2)p(\sigma_u^2)p(\sigma_e^2)$$

Make inferences using MCMC algorithms (Gibbs sampling, acceptance rejection, Metropolis-Hasting)

# Bayesian framework



# EXAMPLE

Create MME and solve the system using residual updates

ID	trait	Diet	SNP1	SNP2
1	20	1	1	2
2	25	1	1	1
3	30	2	0	2
4	35	2	0	1
5	20	3	2	1
6	30	3	2	0

# HOMEWORK

Solve the mixed linear model

$\text{trait} = \mu + \text{age} + \text{Diet} + \text{SNP1} + \text{SNP2} + e$

Using Gauss-Seidel with residual updates, with residual variance = 40 and SNP variance = 3.

ID	trait	Age	Diet	SNP1	SNP2
1	93	25	1	0	0
2	90	30	1	1	0
3	115	35	2	1	2
4	110	20	2	2	2
5	87	22	3	2	1
6	70	29	1	0	1
7	100	31	3	1	2

# HOMEWORK

Solve the mixed linear model

$\text{trait} = \mu + \text{age} + \text{Diet} + \text{Cohort} + e$

Using Gauss-Seidel with residual updates, with residual variance = 40 and cohort variance = 30.

ID	trait	Age	Diet	Cohort
1	93	25	1	1
2	90	30	1	2
3	115	35	2	1
4	110	20	2	2
5	87	22	3	3
6	70	29	1	3
7	100	31	3	3